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Multiscale patterns of rarity in fungi, inferred from fruiting records

Running title: Rarity in European fungi

Abstract

Aim: Characterising the distribution and abundance of organisms is a fundamental part of understanding their population dynamics and development of conservation policies for rare species. It is unknown whether fungi show similar trends to other organisms in their macroecological patterns of abundance and spatial distribution. Here, we investigated fungal abundance-occupancy relationships to determine whether fungi that are common at a local scale tend to be more widely distributed.

Location: UK and Switzerland

Time period: 1950 - 2014

Major taxa studied: Fungi

Methods: We used a local dataset of fruiting records of 2,319 species in the UK, accumulated over 65 years, and one from Switzerland of 319 species, spanning 32 years. Using record number and occurrence as proxies for abundance, in each case we examined the form of species and rank abundance distributions, and compared these with distributions of records in the national databases over the same time. We plotted relationships of local number of records and regional occupancy, and calculated multi-scale indices of rarity for all fungal species.

23 **Results:** There was a remarkable congruence in the patterns found in the UK and
24 Switzerland. Regional assemblages are characterised by many rare species, while few are
25 common (fitting the lognormal distribution). However, at local scales, distributions best
26 fitted a power law, suggesting that habitat availability or dispersal processes may play
27 important roles. Fungi with high local record number are densely distributed nationally, but
28 unlike other organisms, locally rare fungi may also be densely distributed at a wider scale.

29 **Main conclusions:** Fungal fruiting records can be used to infer patterns in fungal
30 distributions. Abundances in local assemblages may be determined by the position of the
31 assemblage in the overall geographic range of each species, dispersal ability and
32 environmental filtering. We advocate the use of multiscale approaches to rarity in future
33 fungal sampling programmes, to provide more reliable information for future conservation
34 policy decisions and fungal biogeography.

35

36 **KEYWORDS**

37 abundance-occupancy, conservation, lognormal, models, mushrooms, rank abundance, fruit
38 bodies

39

40 **1 | INTRODUCTION**

41 Understanding why some species are rarer than others is a fundamental part of community
42 ecology with ramifications in conservation biology, including the management of habitats
43 and natural resources. However, ‘rarity’ itself is a relative concept, in which the abundance
44 or extent of occurrence of a species is defined in relation to that of others (Gaston, 1994).
45 Furthermore, the rarity of a species depends upon the ecological scale at which the taxon is
46 recorded; species may be considered rare at one scale, but common at another (Hartley &

47 Kunin, 2003). Thus, rare species should never be considered in isolation, but as important
48 components of assemblages and ‘hotspots’ of species diversity (Heegaard, Gjerde, &
49 Saetersdal, 2013).

50 The species abundance distribution (SAD) is one of the simplest ways of describing the
51 pattern of relative abundance across the species detected in an assemblage. Fisher, Corbet, &
52 Williams (1943) realised that histograms depicting the frequencies of species abundance
53 show a hollow curve, in which many species consist of a few individuals, while only a few
54 species are abundant. This pattern appears to be universal and thus, in alternative parlance,
55 most species are rare, while only a few are common (McGill et al., 2007). A useful
56 complementary method of describing community assemblies is the rank abundance
57 distribution (RAD), in which species’ abundance is plotted against their rank in abundance
58 (Foster & Dunstan, 2010). RADs can be informative, as they display all the data rather than
59 grouping abundance into ‘bins’, resulting in the masking of some information (McGill et al.,
60 2007).

61 A large body of work exists on fitting models to the hollow curve. However, few models
62 are ever rejected since their prediction ends with stating the nature of the curve, and little
63 attempt has been made to go beyond this and provide explanations for it (McGill et al., 2007).
64 Nevertheless, SADs remain one of the most important tools for describing and understanding
65 community assembly and its management. In particular, they can be useful in informing
66 conservation decisions and determining extinction risk (Matthews & Whittaker, 2015).

67 It is evident that while SADs have been produced for virtually all groups of organisms,
68 their use in fungal ecology is extremely restricted. In a meta-analysis of SADs, only two (of
69 558 distributions) were of fungi, both involving lichens (Ulrich, Ollik, & Ugland, 2010).
70 Lichenised fungi, which are macroscopic organisms, are often treated like plants in ecological
71 studies, but non-lichenised fungi are often studied completely separately. This is at least

72 partly because non-lichenised fungi are invisible for most of their lives as mycelium within
73 soil, wood or some other solid substrate, making enumeration almost impossible (Watkinson,
74 Boddy, & Money, 2015). However, Ascomycetes and Basidiomycetes periodically produce
75 macroscopic fruit bodies which can be enumerated.

76 The development of molecular techniques to detect hidden mycelia is revolutionizing the
77 study of fungal communities, and has provided indications of patterns in global fungal
78 biogeography (Tedersoo et al., 2014). However, SADs and RADs depend on recording the
79 numbers of individuals of each species, which is still not practicable on a large scale using
80 molecular approaches.

81 Based on the meta-analysis of macro-organisms other than fungi, Ulrich et al. (2010)
82 concluded that fully-censused assemblages tend to show SADs that are best described by a
83 lognormal distribution, while assemblages that are incompletely sampled tend to show
84 distributions that are best described by the log series model or a type of power law. These
85 latter distributions still show extreme skew (very many rare species), even when plotted on a
86 log scale. More recently, Dumbrell, Nelson, Helgason, Dytham, & Fitter (2010a) examined
87 the abundance of arbuscular mycorrhizal fungi (AMF) in 33 different communities, testing
88 three different models: the lognormal fitted in 73% of cases, while 27% were described by
89 the broken stick model and none by the geometric series. Unterseher et al. (2011) also found
90 that the lognormal model best described AMF abundance, while Dumbrell et al. (2010b)
91 showed that the lognormal also described AMF abundance well, but the best model fit was a
92 neutral model, the zero sum multinomial. This led to the conclusion that niche differentiation
93 processes are important in structuring the community, as well as neutral processes such as
94 dispersal limitation. These findings contrast with small-scale studies of fungi on leaves,
95 where log series models (indicating incomplete sampling) predominate (Thomas & Shattock,

96 1986; Unterseher et al., 2011), though neutral models have also provided a good fit (Feinstein
97 & Blackwood, 2012).

98 While the majority of fungal studies have taken place at localized scales, those that used
99 broader scales also showed variation in best model fits, either with lognormal or log series
100 (Nielsen, Kjoller, Bruun, Schnoor, & Rosendahl, 2016) or neutral models (Gumiere, Durrer,
101 Bohannan, & Andreote, 2016). Thus, it is unclear whether fungi do or do not show similar
102 patterns of abundance to other organisms, beyond the fact that most species seem to be rare
103 and few are common (Nemergut et al., 2013). In particular, as rarity is scale-dependent, a
104 true depiction can only be achieved by multiscale comparisons (Leroy, Canard, & Ysnel,
105 2013), yet such an approach has never been applied to fungi.

106 The interrelationship of rarity and scale is formalized in another long established pattern in
107 macroecology: the abundance-occupancy relation (i.e. the relation between the local
108 abundance of species and the size of their ranges within a region). A large body of literature,
109 both theoretical and empirical, has shown that this is another universal and positive relation;
110 species that are locally abundant tend to occupy wider ranges, i.e. they are more widespread
111 (Gaston et al., 2000; Borregaard & Rahbek, 2010). However, while a number of
112 comprehensive reviews list the wide variety of taxonomic groups that show such
113 relationships (e.g. Holt, Gaston, & He, 2002), fungi are absent from all such analyses. It is
114 intriguing that the closest approach is that of Roney, Kuparinen & Hutchings (2015), in
115 which lichens were the only taxonomic group in Canada not to show a positive relation,
116 though sample size (n=15 species) was small.

117 Abundance-occupancy relationships are not just important in understanding the structuring
118 forces in local communities, but also for their role in species conservation and habitat
119 management (McGeoch & Latombe, 2016). These relationships, in addition to analysis of
120 SADs and RADs, could provide essential tools for characterization of rare fungi and their

121 conservation, but to date, such an approach is lacking. In general, for the very same reasons
122 of crypsis and sampling outlined above, fungi have lagged behind most taxonomic groups in
123 assessments of rarity and the construction of Red Lists for their protection (Dahlberg,
124 Genney, & Heilmann-Clausen, 2010). However, many species of fungi produce fruiting
125 structures (sporocarps) that can be counted, thereby providing a good proxy for individual
126 abundance, and from which designations of ‘common’ or ‘rare’ can be inferred (Dahlberg &
127 Mueller, 2011). While assigning fruit bodies to different individuals may be problematic at
128 very small scales (< 10 m, Dahlberg & Mueller, 2011), records of occurrence across
129 geographic ranges, calculated as the ‘area of occupancy’ (Gaston, 1994), can provide data
130 that are of similar quality to other taxonomic groups (Truong et al., 2017). The limiting
131 factor is then the distribution and knowledge of recorders, but coordinated large scale surveys
132 have enormous untapped potential to provide information on fungal species abundance and
133 distributions which we currently lack (Molina, Horton, Trappe, & Marcot, 2011; Andrew et
134 al., 2017). Indeed, surveys of fruit bodies are an accepted method for biodiversity
135 assessments, and often reveal species which are undetected by sequencing methods (Runnel,
136 Tamm, & Lohmus, 2015). Furthermore, while offering great promise for the future,
137 molecular methods cannot currently be used to perform macroecological studies of the types
138 reported here. This is because of the problems that exist within sequence databases, due to
139 the high number of unnamed and incorrectly named species, plus primer and other
140 methodological biases which are particularly acute for soil-dwelling species (Khomich et al.,
141 2018).

142 Coordinated databases of the occurrence of fungal fruit bodies have been used successfully
143 to document recent changes in the phenology and spatial distributions of fungi, in response to
144 changing climate (Boddy et al., 2014; Gange et al., 2018). Similar such databases and
145 museum collections have been used in a wide variety of plant and animal studies to examine

146 distributions of species abundances and ranges, to show changes over time, and ultimately to
147 inform conservation policy (Pyke & Ehrlich, 2010). However, with the exception of a couple
148 of notable animal and plant pathogens, fungi are again absent from such analyses.

149 Here, we use two databases of fungal fruiting records, including lichenised fungi,
150 assembled in the UK over the last 65 years: (1) a local data set comprising haphazard
151 collections, with at least weekly frequency from 1950 to 2014, covering an area of
152 approximately 3 000 km², part of which was originally used to document phenological
153 changes (Gange, Gange, Sparks, & Boddy, 2007); and (2) data for the whole of the UK, taken
154 from The Fungal Records Database of Britain and Ireland (FRDBI; www.fieldmycology.net).
155 We also use two data sets from Switzerland; a local study in five plots, each of 300 m² (three
156 10 x 10 m) plots at the La Chanéaz Forest Reserve, comprising weekly fruit body counts
157 from May to December from 1975-2006 (described in Andrew et al. 2018) and data for the
158 whole of Switzerland over the same period (www.swissfungi.ch). A part (28 y) of the local
159 data set was used in a general analysis of varying bin sizes and species abundance
160 distributions (Straatsma & Egli, 2012). To our knowledge, no other local data sets in the
161 world are as comprehensive in their extent and time span as these (Andrew et al. 2017).

162 Our first objective was to examine the SAD and RAD from each of these datasets, using
163 fungal records to calculate abundance on the basis of both the number of records and the
164 occupancy. Our hypothesis was that national scale data would show the classic lognormal
165 distribution, indicating ‘complete sampling’, while the local sets may show different (log
166 series or power law) patterns, suggesting dispersal limitation or niche-related processes
167 (Ulrich et al., 2010). Our second objective was to examine the abundance-occupancy
168 relationships for these fungi, and we hypothesised that these would be positive, given that
169 other microbial abundance patterns such as species-area relationships seem to mirror those of
170 other organisms (Nemergut et al., 2013). Finally, we examined the multiscale patterns of

171 rarity in fungi in the UK and Switzerland, using the approach of Leroy et al. (2013). In
172 general, species that are rare at a local scale also tend to be rare at a wider, regional scale
173 (Freckleton, Gill, Noble & Watkinson, 2005). If the long-established assertion that
174 ‘everything is everywhere, the environment selects’ (O’Malley, 2008) is correct, then fungi
175 should follow a similar pattern. However, fungi are often dispersal-limited (Molina et al.,
176 2011; Peay, Kennedy & Talbot, 2016), so the relation may be far less obvious.

177

178 2 | METHODS

179 2.1 | Composition of the data sets

180 The UK local data set consists of 62,087 occurrence records of 2,319 fungal species,
181 spanning the years 1950 – 2014 (inclusive). A total of 310 observers contributed records
182 from 1,558 different localities, all within a 30 km radius of Salisbury, Wiltshire, UK (51.068°
183 N, 1.795° W), covering an area of 2,828 km². All records and identifications were confirmed
184 by the late E.G. Gange, with problematic species being confirmed by Royal Botanic Gardens,
185 Kew. Localities were visited on a haphazard basis, but each was visited at least once per
186 year. Each record was referenced by a six figure Ordnance Survey grid reference
187 (<https://www.ordnancesurvey.co.uk/>), allowing occupancy to be calculated at a resolution of
188 1 km x 1 km squares (out of a total of 614). Further details on record assembly are provided
189 in Gange et al. (2007). ‘Abundance’ was defined as the total number of records of
190 occurrence (i.e. not the total number of fruit bodies) for each species, and area of occupancy
191 as the total number of 1 km x 1 km squares in which each species was recorded, so as to be
192 comparable with similar previous analyses (Gaston et al., 2000).

193 The full UK national data set spans a wider time scale and also contains records for the
194 island of Ireland. These data are from multiple sources, contributed by individuals, foray

195 lists, scientific societies, herbaria records, and publications of the British Mycological
196 Society. We excluded records collected: (1) before 1950, (2) from outside mainland Britain,
197 and (3) which were missing location data, leaving 1,361,069 separate data points for the
198 2,319 species from 55,882 localities over an area of approximately 209,330 km². These
199 records do not include the local data (above), which have since been incorporated into the
200 national set. Each record was referenced at the resolution of 10 km x 10 km; the standard
201 grid system in the UK. As above, ‘abundance’ was measured as the total number of records
202 for each species, while occupancy was the number of 10 km x 10 km squares in which each
203 species was found. A complete list of all species used is given in Supplementary Table S1.

204 Sampling for the Swiss local data set is described in full in Heegaard et al. (2017) and
205 Andrew et al. (2018), and full species lists are given in the supplementary material of Andrew
206 et al. (2018) and Table S2. Weekly fruit body counts took place from 1975 – 2006, but the
207 10 x 10m plots were only divided into 1 x 1 m contiguous sub plots in 1992. Thus, the
208 description of occupancy is not relevant in this data set, as the scale is too small and the data
209 consists of the total number of records of 319 separate species. A subset of the complete
210 Swiss national dataset was used, to include the same 319 species over the same time span as
211 the local data, producing 97,358 separate records covering an area of 41,285 km².
212 Occupancy at the 10 km x 10 km scale was defined in the same way, using grid references as
213 above.

214

215 2.1 | Data analyses

216 2.1.1 | Local dataset species accumulation curves

217 We first examined the nature of the local data sets by calculating species accumulation curves
218 and estimating the species pool which may exist if all possible species were found. The latter

219 was examined with three methods: Chao 2, first order jackknife, and Bootstrap. These
220 analyses were performed with the package ‘vegan’ in R 3.4.1 (Oksanen et al., 2017; R Core
221 Team, 2017).

222

223 2.1.2 | Species abundance distributions

224 All species abundance distributions were fitted using the R package ‘sads’ (Prado, Miranda,
225 & Chalom, 2016) which uses maximum likelihood methods to fit and compare different
226 models. These were the gamma, lognormal, and Weibull (the three most commonly used
227 continuous distributions), plus the geometric and the negative binomial models. These were
228 fitted with zero truncation, since species with zero records were unknown. We also fitted
229 Fisher’s log series, and three associated power law functions, Pareto, power, and power bend
230 (implementation of the latter two not including zeroes); the log series being a special case of
231 the power bend. We also fitted MacArthur’s broken stick model and the Poisson lognormal
232 model; the latter describes species’ abundances in a Poisson sample from an underlying log
233 normal assemblage. Finally, we examined the fit of two neutral models, the metacommunity
234 Zero-sum multinomial distribution, which is thought to describe a community under random
235 drift, and the Volkov model, thought to describe a community under neutral drift, with
236 immigration. References to the use of all models are provided in Prado et al. (2016).

237

238 2.1.3 | Rank abundance distributions

239 Rank abundance distributions were fitted using the ‘radfit’ command in the R package
240 ‘vegan’ in a similar manner. In this case, we used the niche preemption model (also termed
241 the Geometric Series), lognormal, broken stick, and two discrete power law distributions: the
242 Zipf and Zipf-Mandelbrot. In all cases, Akaike Information Criteria (AIC) values were

243 computed to determine which model(s) provided the best fit to the data (Burnham &
244 Anderson, 2002). Those produced for each model were compared using the delta AIC (Δ_i). A
245 Δ_i value < 2 indicates interaction between models, values $3 < \Delta_i < 9$ indicate little interaction
246 and $\Delta_i > 10$ no interaction (Burnham & Anderson, 2004).

247

248 2.1.4 | Relationships between abundance at local and national scales

249 Relationships between the number of records at local and national scales were examined with
250 a Generalised Additive Model (GAM) procedure, using the ‘mgcv’ package in R. To
251 examine abundance-occupancy relationships, we followed Holt & Gaston (2003) and
252 Zuckerberg, Porter & Corwin (2009) in using the logit transformation for occupancy data and
253 the log transformation for the number of records, expressed as the total number for each
254 species over the 65 and 32 year periods. As Zuckerberg et al. (2009) comment, phylogenetic
255 approaches are rarely required in abundance occupancy analyses, since closely related species
256 can vary greatly in their distributions and population sizes. As a check, we examined some of
257 the larger genera represented in the UK databases, *Cortinarius* (88 species), *Russula* (72
258 species), *Mycena* (59 species), *Lactarius* (44 species), and *Entoloma* (44 species). For these
259 genera, the record data at the local scale showed huge variation, with a ratio between the most
260 and least numerous species of 34, 100, 104, 64, and 36 times respectively. Thus, even at the
261 local scale, the variation in the number of records of species within genera is so great that
262 controlling for phylogenetic relatedness is unwarranted, and we conducted an ordinary least
263 squares regression procedure, with occupancy as the dependent variable (Holt & Gaston,
264 2003).

265

266 2.1.5 | Multiscale rarity

267 If non-linear relationships between abundance at different scales are found, then a multiscale
268 approach is justified (Leroy et al., 2013). As this was so, we calculated two-scale (local and
269 national) rarity weights for all species in the UK and Switzerland, using the R package
270 ‘rarity’, (Leroy et al., 2013). These were scaled between 0 and 1, with a greater degree of
271 rarity being indicated by a higher value of the index. Calculation of these weights involves
272 an appropriate weighting method and the use of a rarity cut-off value. To achieve this, we
273 used Gaston’s quartile definition - that rare species are those 25% with the lowest number of
274 records or occurrence (Gaston, 1994; Leroy et al., 2013).

275

276 3 | RESULTS

277

278 3.1 | Characteristics of the local data sets

279 There was no trend over time in the number of collections, or forays per year in the UK
280 (Figure 1a, $F_{1,63} = 3.51$, $P > 0.05$). It should be noted that databases of this type do not record
281 forays when no species were found. Low numbers in years such as 1976, 1990 and 2011
282 represent poor years for fruiting, caused by lack of rainfall.

283 The total number of species (the ‘species pool’) estimated from the UK species
284 accumulation curve (Figure 1b) varied depending on the method, with the Chao estimating
285 $3,526 \pm 96.9$, jackknife $3,313 \pm 141.4$, and bootstrap $2,852 \pm 80.8$. Therefore, these values
286 suggested that between 68% and 84% of the total ‘available’ species were detected in the 65
287 years of recording.

288 The total species pool estimates for the Swiss local data were more consistent, with the
289 Chao estimating 388.75 ± 21.8 , jackknife 388.75 ± 17.4 and bootstrap 351.19 ± 10.7 . These

290 values suggested that 82 – 90% of the total species were recorded in the 32 years of sampling.
291 The species accumulation curve is shown in Figure 1c.

292

293 3.2 | Species and rank abundance distributions

294 The species abundance distribution for UK national scale records was best fitted by the
295 lognormal distribution (AIC = 33 305.1) (Figure 2a). No other models provided a good fit to
296 the data (Supplementary Table S3), with the next best fit provided by the Poisson lognormal
297 ($\Delta_i = 18.0$, $P < 0.001$). There was no interaction between models ($18 < \Delta_i < 3592$). The SAD
298 for national scale occupancy data was also fitted best by the lognormal (AIC = 29 093.6)
299 (Figure 2b), this being better than the Weibull model ($\Delta_i = 9.2$, $P < 0.01$). The rank
300 abundance distribution for national record number was also best fitted by the lognormal (AIC
301 = 84 952) (Figure 2c), but that for occupancy data was best fitted by the Preemption model
302 (Geometric series) (AIC = 24 202) (Figure 2d). With both RADs, fitting of the Zipf-
303 Mandelbrot failed, as its estimation is difficult (Oksanen et al., 2017).

304 The Swiss national record number data was also best fitted by the lognormal distribution
305 (AIC = 3997.6) though the Weibull and Broken stick also provided a good fit (Table S3,
306 Figure S1). The SAD for Swiss national scale occupancy data was also best fitted by the
307 lognormal (AIC = 3248.4). The rank abundance distribution for Swiss national record data
308 was best fitted by the Broken stick, while occupancy data were best fitted by the Geometric
309 series (Table S3).

310 The species abundance distribution for UK local record numbers was best fitted by the
311 Pareto distribution (AIC = 1 5594.8) (Figure 3a), as was that for the Swiss local record
312 numbers (Table S3, Figure S1). Meanwhile, the abundance distribution for UK local
313 occupancy was also best fitted by the Pareto distribution (AIC = 13367.4) (Figure 3b) (Table

314 S3). In both cases, all other models differed from the Pareto ($P < 0.001$). Rank abundance
315 fits for UK local data followed an identical pattern to national data (Table S3), fitting a
316 lognormal and Geometric series respectively (Figure 3c,d). Swiss local record data differed
317 from the national data, wherein the lognormal provided the best fit (Table S3).

318 3.3 | Abundance-occupancy relationships

319 Species with the greatest number of records at the UK national scale also had the greatest
320 number at the local scale (Figure 4a). These data were best fitted by a non-linear model ($R^2 =$
321 58.8% , $P < 0.001$), rather than a linear relation ($R^2 = 47\%$), because species with very few
322 records (i.e. were 'rare') locally may have few or very many records nationally, illustrated by
323 the flat bottom to the graph. However, there appeared to be a tipping point, with very
324 common species (more than 1,000 records nationally) showing a linear relation with the local
325 number of records. A similar relation was seen for occupancy data; species that are most
326 densely distributed at the national scale are also so at the local scale (Figure 4b), but locally
327 sparsely distributed species may be sparsely or densely distributed nationally. A non-linear
328 relation ($R^2 = 57\%$, $P < 0.001$) was also seen in these data, rather than a linear one ($R^2 =$
329 44.4%), with a tipping point of occurrence in about 200 of the national 10 km x 10 km
330 squares.

331 Swiss record data followed a very similar pattern (Figure S2) and were best fitted by a
332 non-linear model ($R^2 = 21.1\%$, $P < 0.001$), rather than a linear relation ($R^2 = 15.1\%$). There
333 again appeared to be a tipping pint, with very common species (more than 350 records
334 nationally) showing a linear relation with local record number (Figure S2).

335 The abundance-occupancy relationships showed a remarkable similarity in the two
336 countries (Figure 5). Although both relationships are significant, they are relatively weak
337 (UK: $R^2 = 43.9\%$; Swiss: $R^2 = 13.8\%$), as species that are rare (least abundant) on a local

338 scale can be sparsely or densely distributed on a national scale. In contrast, species with high
339 local record numbers tend to be densely distributed nationally. Species with high local record
340 numbers but sparse distributions nationally are absent in both data sets.

341

342 3.3 | Multiscale indices of rarity

343 In the UK, most, but not all, species that are rare on a national scale are also rare on a local
344 scale ($R^2 = 65.2\%$, $P < 0.001$) (Figure 6a), and the significant relation is clearly driven by the
345 preponderance of data with high rarity indices. Likewise, those which are sparsely
346 distributed nationally also tend to be sparsely distributed on a local scale (Figure 6b), though
347 the relation is considerably weaker ($R^2 = 49.8\%$, $P < 0.001$) and the pattern more diffuse.
348 The pattern is even more accentuated in Swiss record data (Figure S3) ($R^2 = 12.3\%$, $P <$
349 0.001), with the majority of species being rare at the local scale. In these data, there were no
350 species that were locally common but nationally rare.

351 The proportions of species falling into the four possible categories of rarity generated by
352 the two scale approach are given in Table 1. For both data sets, the vast majority of species
353 can be considered rare at both spatial scales. Only between 1% (based on number of records
354 in UK and Switzerland) and 6% (based on occupancy in UK) of species could be considered
355 common at both spatial scales (Table 1).

356

357 **Discussion**

358 This is the first macroecological study of rarity patterns in fungi. By using four
359 comprehensive data sets, we have shown that while some similarities exist in patterns of
360 abundance between fungi and other organisms, there are also noticeable differences. Both
361 UK and Swiss data showed the classic hollow curve of species abundance, but while the

362 national data were best fitted by a lognormal model, the local data sets were uniquely fitted
363 by a Pareto distribution (with very many rare species). Abundance-occupancy relationships
364 of fungi were positive, and similar in the UK and Switzerland, but showed a different pattern
365 to those of other organisms.

366 Perhaps the most obvious conclusion from these data is that irrespective of the approach
367 taken, most species of fungi are rare, while only a few can be considered common. Both
368 number of records and occupancy at the national scales were best described by the lognormal
369 distribution, as with marine and soil bacteria (Fuhrman, 2009; Ferrenberg et al., 2013;
370 Nemergut et al., 2013) and most macroorganisms (Ulrich et al., 2010). There has been much
371 debate on whether lognormal SADs are the product of sampling artefacts, model fitting, the
372 influence of environmental variables, or the apportionment of niches between species
373 (Williamson and Gaston, 2005; McGill et al., 2007). Furthermore, we must not forget that
374 the lognormal is a purely statistical distribution. However, despite the potential for artefacts,
375 the finding that lognormal distributions tend to arise in assemblages that are completely
376 sampled (Ulrich et al., 2010), strongly suggests that there are good biological reasons for
377 such patterns, including niche processes, competition and dispersal.

378 The dictum for microbes that ‘everything is everywhere, but the environment selects’,
379 (reviewed by O’Malley (2008)), has been challenged in many microbial studies (Martiny et
380 al., 2007). Fungal species certainly exhibit biogeographical patterns at continental or smaller
381 scales (Taylor et al. 2006; Tedersoo 2017). In general, everything is not everywhere, but
382 environmental filtering certainly plays a role in determining fungal distributions, with
383 dispersal limitation being proposed as one of the main drivers (Peay et al., 2016). We found
384 variation in the occupancy within the geographic range of fungi of two or three magnitudes at
385 both national scales and at the local scale in both countries, clearly supporting the fact that
386 not all species occur everywhere within their range. Furthermore, relatively low amounts of

387 variation were explained by our analyses of local vs. national record number and occupancy,
388 and also with the multiscale analysis of rarity indices, suggesting that environmental filtering,
389 most likely manifest as habitat availability, plays an important role. This was particularly
390 true in Switzerland, where landscape structure is more heterogeneous than the UK (Hofer,
391 Wagner, Herzog & Edwards, 2008).

392 A recent meta-analysis suggests that few fungi may be habitat generalists (Meiser, Bálint
393 & Schmitt, 2014) and so the availability of habitats in the local environment determines
394 species occurrence (Kivlin, Winston, Goulden, & Treseder, 2014). This fact is further
395 supported by the two local data sets displaying a fit to a power law, rather than the lognormal,
396 upholding our first hypothesis and suggestive of niche-related processes (i.e. habitat
397 availability) playing a role (Ulrich et al., 2010). Both these data sets are free from many
398 forms of bias that influence such analyses (Gange, Gange, Mohammad, & Boddy, 2012;
399 Lavoie, 2013), particularly as they were coordinated regular surveys over long time periods,
400 that were not influenced by citizen scientists recording certain species (e.g. edible fungi) or
401 searching in known localities (Geldmann et al., 2016). However, such model fits may still be
402 indicative of incomplete or imperfect sampling (McGill, 2003; Ulrich et al., 2010),
403 particularly as neither species accumulation curve was asymptotic. A further complication is
404 that spatial distributions of saprotrophic and ectomycorrhizal fungi in the UK have changed
405 in recent years, correlated with changes in temperature and rainfall (Gange et al., 2018).

406 This is the first time that a Pareto distribution has been fitted to fungal data sets, though
407 previous authors generally examined a limited range of models (e.g. Dumbrell et al., 2010a;
408 Unterseher et al., 2011; Gumiere et al., 2016). Pareto distributions are characterised by the
409 presence of many species of low abundance (often singletons) and far fewer of high
410 abundance, producing a long 'thin' tail to the SAD. Such distributions are widely reported in
411 physics, computer science, economics and social sciences as well as biology (Newman,

412 2005). Our data are remarkably similar to those of many invertebrate distributions, with 28%
413 singletons in the UK local fungal data, and 9% in the Swiss data, compared with 29-32% for
414 tropical arthropods (Coddington, Agnarsson, Miller, Kuntner, & Hormiga, 2009) and 17% for
415 temperate spiders (Leroy et al., 2013). However, this is considerably less than the 53%
416 singletons recorded in a short term (1 year) survey of fruit body abundance of temperate
417 forest fungi in northern Spain (Abrego & Salcedo, 2014).

418 It is most likely that dispersal ability and niche availability influences the numbers of very
419 rare fungal species (Unterseher et al., 2011). In British lichens, species with high
420 colonization ability (dispersal) occupy larger geographic ranges (Leger & Forister, 2009) and
421 this may be true for fungi generally (Cox, Newsham, Bol, Dungait, & Robinson, 2016), so
422 limited dispersal ability may explain why many species are rare (Molina et al., 2011).
423 Meanwhile, since the number of fungi in an assemblage may be functionally related to the
424 number of plants (Hawksworth (1991), (though see also Tedersoo et al. (2014) for a
425 contradiction with soil fungi) and assemblages of plants, even when completely censused,
426 appear to follow power law distributions (Ulrich et al., 2010), the availability of plant
427 substrates (niches) may be a major factor influencing fungal distributions.

428 Fungal species that were rare (measured as either records or occupancy) on a national
429 scale were also rare on a local scale. The national number of records explained about 47-
430 58% of the variation in local record number, similar to the 31% for phytoplankton and 58%
431 for aquatic bacteria (Ostman et al., 2010). This conclusion was also reinforced by the
432 multiscale analyses; here the majority of species were rare at both scales. However, this
433 analysis showed that it is quite possible for all parts of the graphical spectrum to be occupied,
434 further refuting the ‘everything is everywhere...’ idea and upholding our original hypothesis.
435 Calculation of the rarity weights for each species is an important aspect of our analysis, as it
436 gives a numerical index for each species, which is far more informative than simple

437 categories (Leroy et al., 2013). In future, such indices could help with the construction of
438 Red Lists, provided that the quantitative information can be related to IUCN criteria adapted
439 for fungi (Dahlberg & Mueller, 2011).

440 Our study is also the first to find good evidence for a positive abundance-occupancy
441 relation for fungi, upholding our second hypothesis. Comparison with previous studies is
442 hampered by authors using different measures of occupancy and abundance or by treating
443 abundance as the dependent variable. However, the form of the fungal relation is different to
444 that of other analyses which have used the same approach (logit and log, with occupancy on
445 the y axis) (e.g. Holt & Gaston, 2003; Zuckerberg et al., 2009), which show clear linear
446 relationships, with R^2 values between 60 and 90%. The fungal relationships, although very
447 similar in the two countries, were much more diffuse and bear some similarity to that of
448 marine bacteria (Amend et al., 2013), but not to intestinal bacteria (Green, Fisher, McLellan,
449 Sogin, & Shanks, 2016). Our data are similar to all others in that there were no locally
450 abundant species with sparse national distributions, but the critical difference is that there
451 were many locally rare species which are densely distributed at a wider scale.

452 Eight possible mechanisms which might account for a positive abundance-occupancy
453 relation have been proposed (Gaston, Blackburn, & Lawton, 1997; Gaston et al., 2000). The
454 first pertains to sampling bias, which may result from low sampling intensity at small spatial
455 scales. We do not believe that poor sampling in the local data sets has contributed to the
456 observed relation, as there was no trend in 'foray effort' over time in the UK and
457 standardized weekly counts were conducted in Switzerland. Furthermore, the UK local data
458 comprised 6,868 separate forays and each involved the collection and identification of every
459 fungus seen, while the Swiss data comprised 992 separate sampling occasions. Citizen
460 science data, while not without its faults, can be used for macroecological analyses, if
461 collated properly, and many previous analyses of this type have used such data (Dickinson,

462 Zuckerberg & Bontner, 2010). Indeed, patterns of phenology in these national and local data
463 sets are remarkably similar (Andrew et al., 2018), suggesting that the local data sets were not
464 biased towards or against certain species. It is possible that model fits might change if all of
465 the local species estimated to be in the pool were found. However, given that all of the
466 ‘missing’ species must be represented by very low numbers of records, the most likely
467 outcome would be that the model fit will remain the same, while the thin tail of the SAD
468 would be extended.

469 Phylogenetic relatedness has been suggested as a second possible, but unlikely, influence
470 on the relation (Gaston et al., 1997) and is also unlikely in our study due to the enormous
471 variation in abundance and range within almost all genera. A third proposal is that the
472 position in the overall geographic range of a species is important in determining its local
473 abundance; species at the edge of their overall range generally have lower abundance. It is
474 interesting that a prediction arising from this hypothesis is that widespread species may show
475 high or low local density, while species with restricted geographic ranges only have low
476 density (Gaston et al., 1997). This would give a triangular abundance-occupancy relation, as
477 found in this study. Given the differences in UK climate from N-S and W-E, and the
478 sensitivity of fungi to climate (Boddy et al., 2014; Gange et al., 2018), it is plausible that
479 locally rare species might vary greatly in their geographic range, depending on what part of
480 the overall range the local area occupies. The other five hypotheses, based on resource use,
481 resource availability, habitat selection, metapopulation dynamics or vital rates seem far less
482 likely to apply to fungi.

483 Overall, using four extensive data sets, we have shown that fungi exhibit some markedly
484 different macroecological patterns to other organisms. In particular, the abundance-
485 occupancy relationships for fungi are very different, and suggest that the forces that
486 determine commonness and rarity in other organisms are different for fungi. It is clear that

487 after either 32 or 65 y of intensive sampling in the two localities, many fungal species
488 remained undetected and that the vast majority of species could be considered as rare.

489

490 AUTHOR CONTRIBUTIONS

491 ACG and EGG designed the study. EGG collected the local data in the UK, which was
492 analysed by ACG, LPA and AN. SE collected and collated the local data in Switzerland, B
493 S-I coordinated and georeferenced the Swiss national data, CA harmonised, formatted and
494 prepared all the Swiss data, which were analysed by ACG. ACG, LPA, CA, and LB wrote
495 the paper, with contributions from all other authors.

496

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684 **Table 1** Percentage of species in different categories of rarity, as measured by the number of
 685 records and occupancy (in parenthesis for UK data only).
 686

UK, n= 2,319		National	
		Rare	Not rare
Local	Rare	96.2 (85.3)	1.6 (7.5)
	Not rare	0.6 (1.1)	1.6 (6.1)

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Switzerland, n=319		National	
		Rare	Not rare
Local	Rare	87.3	10.9
	Not rare	0	1.8

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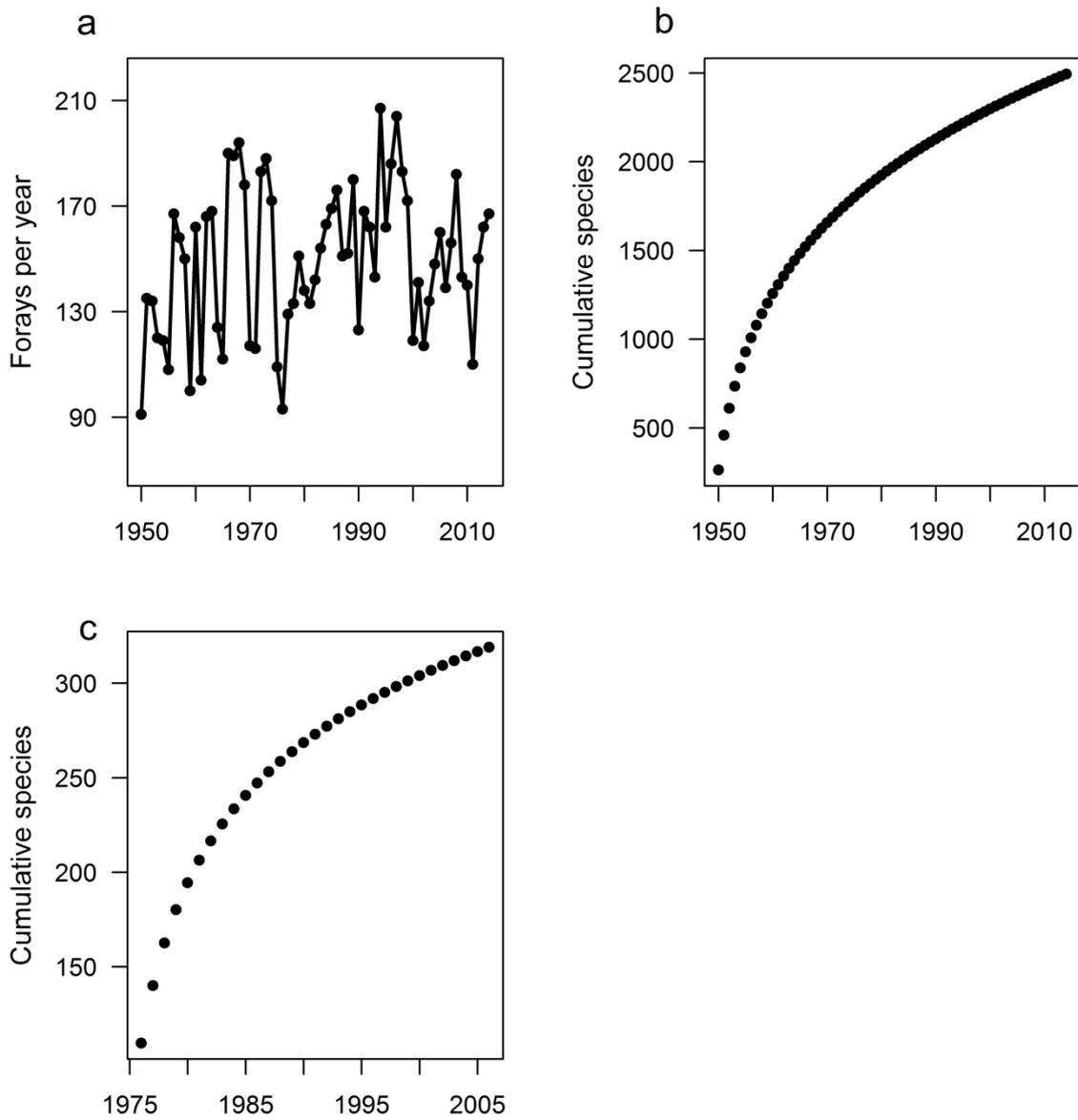
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694 **Figure 1** Temporal characteristics of the local datasets. (a) The number of annual forays in
695 the UK, (b) the accumulated total number of species observed over 65 years in the UK and
696 (c) the accumulated total number of species observed over 32 years in Switzerland.



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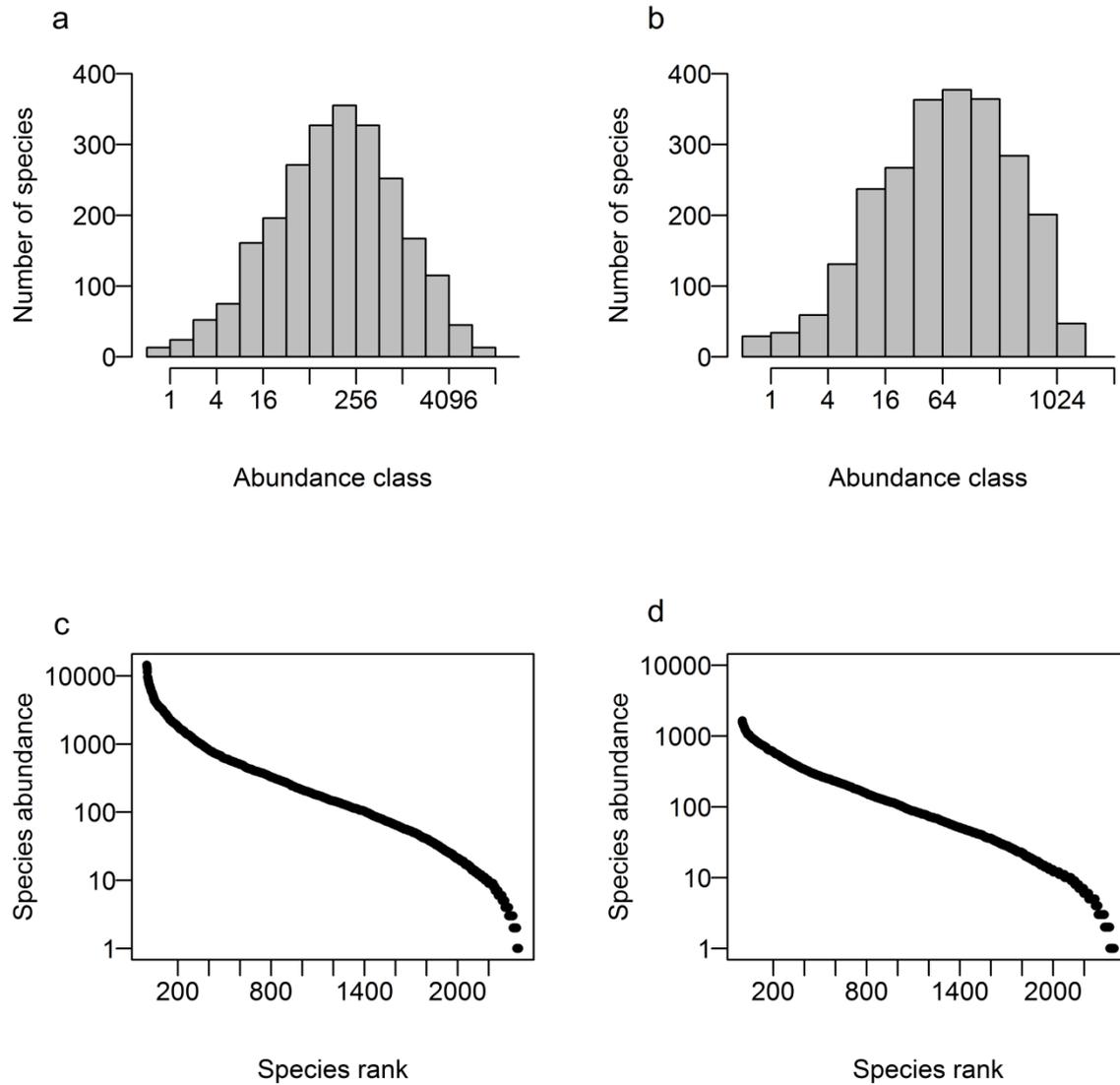
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702 **Figure 2** Frequency distributions for species abundance (SAD) and rank abundance
703 (RAD) of the UK national data set based on (a,c) number of records and (b,d) occupancy.



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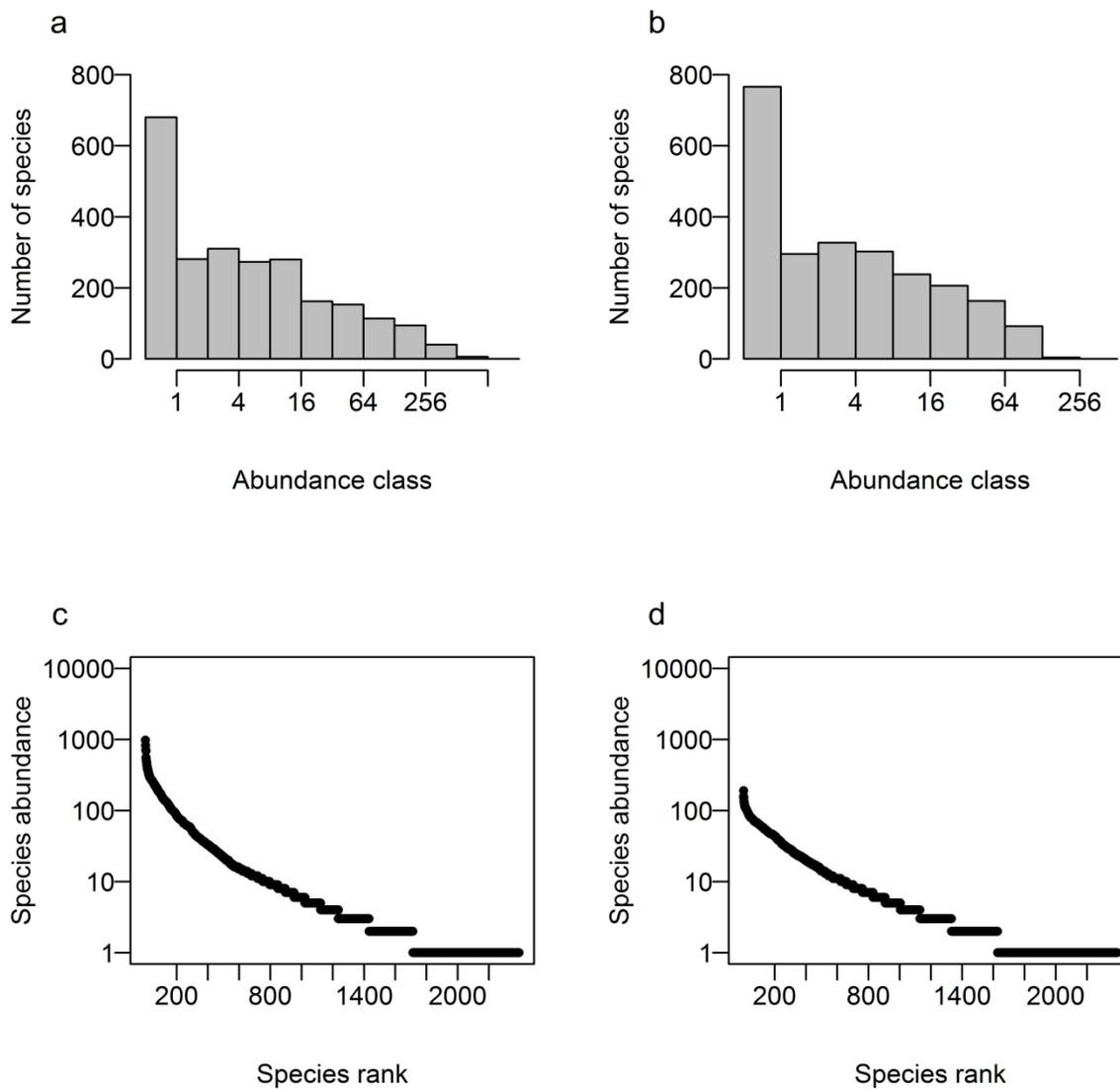
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710 **Figure 3** . Frequency distributions for species abundance (SAD) and rank abundance
711 (RAD) of the UK local data set based on (a,c) number of records and (b,d) occupancy.

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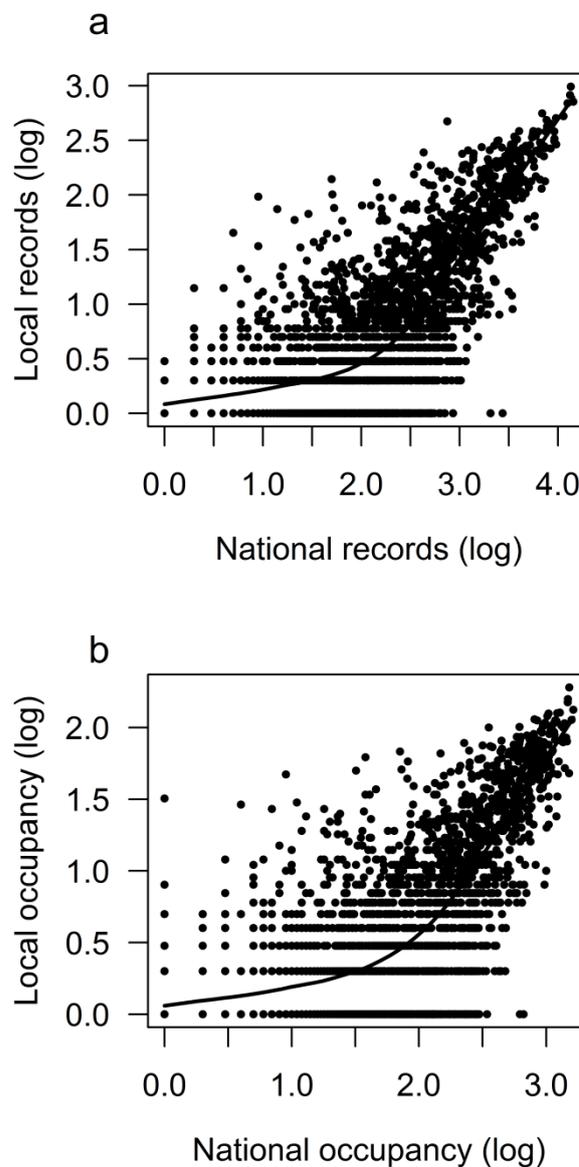
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717 **Figure 4** Relations between UK local and national scale data. (a) Relationship between
718 number of local and national records. (b) Relationship between local and national occupancy.
719 Species with the most records nationally also have most records at a local scale (a), while
720 those which are most densely distributed nationally are also so at a local scale, measured by
721 counts in 10 km and 1 km grids respectively (b). The fitted lines are from a locally weighted
722 regression (loess) procedure.



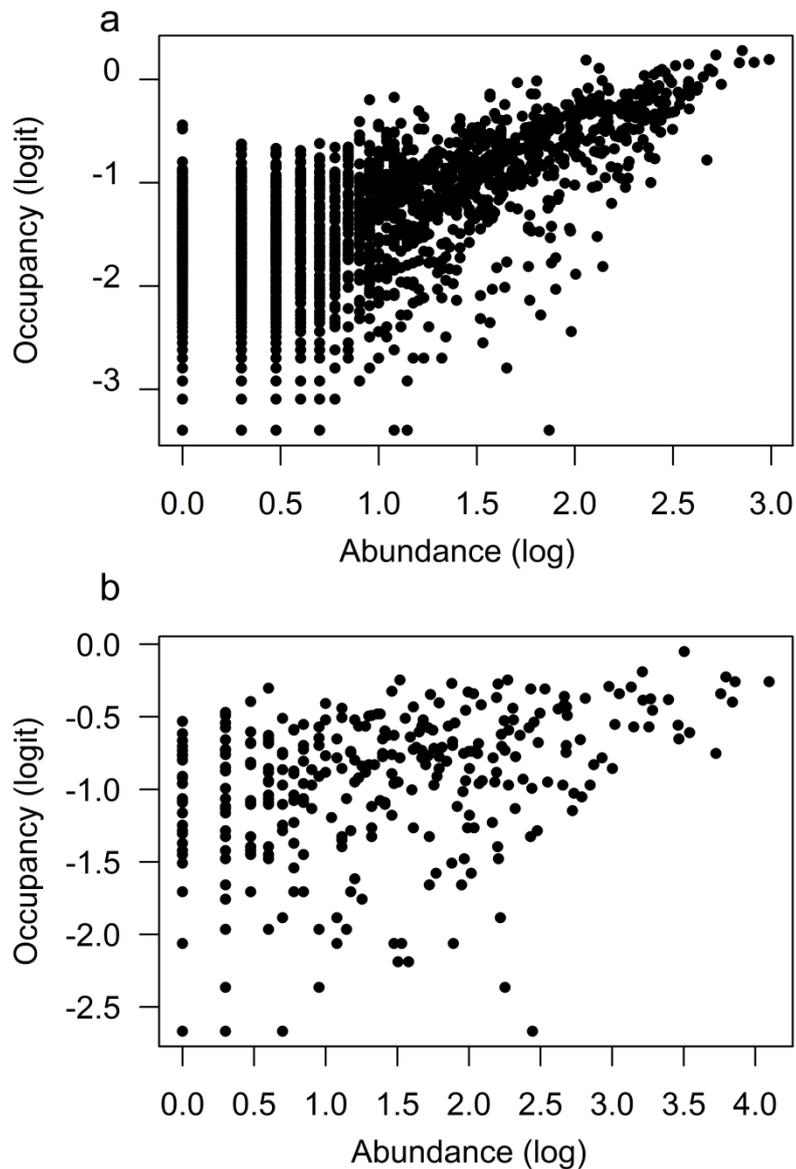
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726 **Figure 5** The abundance-occupancy relations for (a) British and (b) Swiss fungi. Species
727 with high numbers of records locally tend to be densely distributed at a wider scale.
728 However, species with low local record number may be sparsely or densely distributed at a
729 wider scale.

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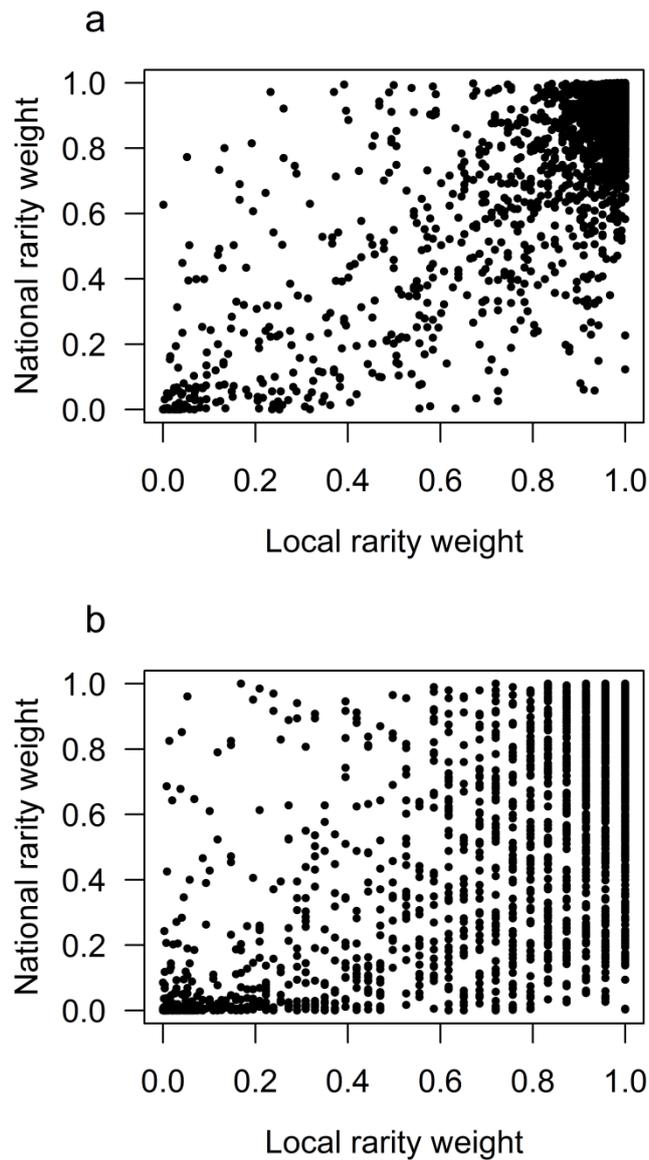


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734 **Figure 6** Multiscale patterns of rarity in British fungi. Using (a) number of records or (b)
735 occupancy, species can be rare at one scale, but common at another and *vice versa*. The
736 majority of species are rare at both scales, using data of the number of records, but the pattern
737 for occupancy is different, suggesting that species with low occupancy locally may be
738 sparsely or densely distributed at a larger scale.



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- 743 **Figure S1.** SAD and RAD for Swiss national data
- 744 **Figure S2.** Relation between Swiss local and national scale data
- 745 **Figure S3.** Multiscale patterns of rarity in Swiss fungi
- 746 **Table S1.** Complete species list for the UK
- 747 **Table S2.** Complete species list for Switzerland
- 748 **Table S3.** Results of model fits
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